#### Lecture Note of STA250

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## **1** Python

- 1. For python 2, the most used version is 2.6 and 2.7
- 2. For python 3, the most used version is 3.2 and 3.3
- 3. Python is good for text processing
- 4. Two ways to run python.
  i, under the environment of python, use the interactive mode.
  ii, Executive python scrip.
- 5. For dealing with string: baz="Class" bar="Hello" foo2=bar+" "+baz
- 6. Indentation is important in python:
- 7. python uses zero indexing.
- 8. Format can only be converted by following order: Str  $\rightarrow$  Float  $\rightarrow$  int

### 2 Bayes

Bayes recap: Assume  $X_1, ..., X_n$  follow i.i.d.  $N(\mu, \sigma^2), \sigma^2$  known and  $\mu \sim N(\mu_0, \sigma_0^2)$ , then

$$\mu | \vec{X} \sim N(\frac{\frac{\mu_0}{\sigma_0^2} + \frac{\sum x_i}{\sigma^2}}{\frac{1}{\sigma_0^2} + \frac{n}{\sigma^2}})$$

Assume that we have general likelihood  $p(\vec{X}|\theta)$  and prior  $p(\theta)$ , thus the posterior will be proportional to:

$$p(\theta | \vec{X}) \propto p(\theta) p(\vec{X} | \theta)$$

If  $p(\theta | \vec{X})$  has no closed form, use MCMC.

### **3** Big Data

For big data, applying MCMC directly to the full dataset is not computationally feasible.

Idea: "chunk" full dataset into a series of smaller datasets, sample from posterior of each smaller dataset and combine results.

Assume that  $X'_i s$  are conditionally independent given  $\theta$ 

$$\prod_{j=1}^{s} p(\theta)^{1/s} p(\vec{X}_{j}|\theta)$$

where  $\vec{X}_j$  is the j-th subset of data and s is the total number of subsets.



where  $\sum q_j = 1, q_j > 0 \ \forall j \text{ and } \cup_j \vec{X}_j = \vec{X}, \cap_j \vec{X}_j = \emptyset$ . Aside,

$$\mu \sim N(\mu_0, \sigma_0^2)$$
$$p(\mu)^{q_j} \propto \exp\{-\frac{q_j}{2\sigma_0^2}(\mu - \mu_0^2)\}$$

# 4 Consensus Monte Carlo

Assume that sample  $\theta_j^{(t)}$  from model  $m_j$  (i.e.  $p_{m_j}(\theta | \vec{X}_j) \propto \pi(\theta)^{q_j} p(\vec{X}_j | \theta)$ ). Combine to

$$\theta^{(t)} = \frac{\sum_{j=1}^{s} w_j \theta_j^{(t)}}{\sum_{j=1}^{s} w_j}$$

where  $w_j > 0$  are known weight. For weights, it turns out that a good good choice is:  $w_j = var^{-1}(\theta | \vec{X}_j)$ .

Example. If  $X_1, ..., X_n | \mu$  follow i.i.d.  $N(\mu, \sigma^2)$  and  $\mu \sim N(\mu_0, \sigma_0^2)$  and chunk data into  $\vec{X}_j = \{X_{j_1}, X_{j_2}, ..., X_{j_k}\}$ . Where  $\{j_1, ..., j_k\}$  are indices corresponding to the subset *j*. Model *j*:

$$X_{j_1}, X_{j_2}, ..., X_{j_k} | \mu \sim_{iid} N(\mu, \sigma^2)$$
  
 $\mu \sim N(\mu_0, \frac{\sigma_0^2}{q_j})$ 

Posterior for Model j:

$$\mu_j | \vec{X}_j \sim N\left(\frac{\frac{q_j \mu_0}{\sigma_0^2} + \frac{\sum x_j}{\sigma^2}}{\frac{q_j}{\sigma_0^2} + \frac{k_j}{\sigma^2}}, \frac{1}{\frac{q_j}{\sigma_0^2} + \frac{k_j}{\sigma^2}}\right)$$

Combine  $\mu = \frac{\sum w_j \mu_j}{\sum w_j}$ ,  $w_j = \frac{\varepsilon_j}{\sigma_0^2} + \frac{k_j}{\sigma^2}$  and  $\sum w_j = \frac{1}{\sigma_0^2} + \frac{n}{\sigma^2}$ . Then, we get

$$Var(\mu) = \frac{1}{\left(\frac{1}{\sigma_0^2} + \frac{n}{\sigma^2}\right)^2} \sum_j w_j^2 Var(\mu_j)$$
$$= \frac{1}{\frac{1}{\sigma_0^2} + \frac{n}{\sigma^2}}$$

$$E(\mu) = \frac{\sum_{j} w_{j} E[\mu_{j}]}{\sum w_{j}} = \frac{\sum_{j} \frac{q_{j}\mu_{0}}{\sigma_{0}^{2}} + \frac{\sum \sum x_{k_{j}}}{\sigma^{2}}}{\frac{1}{\sigma_{0}^{2}} + \frac{n}{\sigma^{2}}}$$
$$= \frac{\frac{\mu_{0}}{\sigma_{0}^{2}} + \frac{\sum x_{j}}{\sigma^{2}}}{\frac{1}{\sigma_{0}^{2}} + \frac{n}{\sigma^{2}}}$$

Here, evenly chunk data is suggested.

Suppose n is large and the sub datasets are also large then we can get the posterior approximates

normal distribution. (MC will still work well)